## Ensemble Filtering in the Presence of Nonlinearity and Non-Gaussianity


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## Preliminaries

## Notation

$\triangleright$ follow Ide et al. (1997) generally, except:
$\ldots \operatorname{dim}(\mathbf{x})=N_{x}, \operatorname{dim}(\mathbf{y})=N_{y}$
$\ldots$ subscript $j: k$ indicates times $t_{j}, t_{j+1}, \ldots, t_{k}$,
... superscripts index ensemble members, or iterations
$\triangleright \sim$ means "distributed as," e.g. $x \sim N(0,1)$
$\triangleright$ state evolution: $\mathbf{x}_{k}=M\left(\mathbf{x}_{k-1}\right)+\eta_{k}$
$\triangleright$ observations: $\mathbf{y}_{k}=H\left(\mathbf{x}_{k}\right)+\epsilon_{k}$

## Basic Facts

1. Conditional pdf $p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)$ is the answer
$\triangleright$ summarizes everything that can be known about state
$\triangleright$ calculate sequentially, via Bayes rule,

$$
p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)=p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}\right) p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k-1}\right) / p\left(\mathbf{y}_{1: k}\right)
$$

$\triangleright$ algorithms that do not produce $p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)$ cannot be fully optimal

## Basic Facts (cont.)

2. Linear, Gaussian systems are relatively easy
$\triangleright p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)$ is Gaussian and thus determined by its mean and covariance
$\triangleright$ posterior (analysis) mean is linear in prior (background) mean and observations
$\triangleright$ no need to choose between posterior mean (min variance) and posterior mode (max likelihood) as "best" estimate; they are equal.
$\triangleright$ 4D-Var and Kalman filter (KF) agree; so does ensemble KF (EnKF) up to sampling error.

## Basic Facts (cont.)

3. High-dimensional pdfs are hard
$\triangleright p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)$ is a continuous fn of $N_{x}$ variables. Direct approaches not feasible; discretization with $n$ points per variable requires $n^{N_{x}}$ d.o.f.
$\triangleright$ they are extraordinarily diffuse

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Consider $\mathbf{x} \sim N(0, \mathbf{I})$.
1 dimension: points with $p(\mathbf{x})$ less than 0.01 of max account for less than $1 \%$ of mass of pdf.
10 dimensions: they account for about $1 / 2$.

## Outline

Nonlinearity and the ensemble Kalman filter (EnKF)
$\triangleright$ Relation to the BLUE
$\triangleright$ Iterative schemes

## Particle filters

$\triangleright$ Required $N_{e}$ grows exponentially w/ "problem size"
$\triangleright$ Importance sampling and the optimal proposal density

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$\triangleright$ Importance sampling and the optimal proposal density
Not a comprehensive review!

## The Best Linear Unbiased Estimator (BLUE)

Desire an estimate of $\mathbf{x}$ given observation $\mathbf{y}=H(\mathbf{x})+\epsilon$
$\triangleright$ Consider linear estimators, $\hat{\mathbf{x}}=\mathbf{A y}+\mathbf{b}$
$\triangleright$ Which $\mathbf{A}$ and $\mathbf{b}$ minimize $E\left(|\mathbf{x}-\hat{\mathbf{x}}|^{2}\right)$ ?

## The BLUE (cont.)

The BLUE is the answer
$\triangleright$ Let $\overline{\mathbf{x}}=E(\mathbf{x})$ and $\overline{\mathbf{y}}=E(\mathbf{y})=E(H(\mathbf{x}))$
$\triangleright$ Then BLUE is given by (e.g. Anderson and Moore 1979)

$$
\hat{\mathbf{x}}=\overline{\mathbf{x}}+\mathbf{K}(\mathbf{y}-\overline{\mathbf{y}}), \quad \mathbf{K}=\operatorname{cov}(\mathbf{x}, \mathbf{y}) \operatorname{cov}(\mathbf{y})^{-1}
$$

$\triangleright$ Only need 1st and 2nd moments; no requirement that $\mathbf{x}, \epsilon$ are Gaussian or $H$ is linear

Useful benchmark for nonlinear, non-Gaussian systems
$\triangleright \ldots$ though $E(\mathbf{x} \mid \mathbf{y})$ has smaller expected squared error

## Relation of EnKF to the BLUE

Start with $\mathbf{x}^{f}$ drawn from $p(\mathbf{x})$
EnKF update specifies a random, linear fn of $\mathbf{x}^{f}$ and $\mathbf{y}$
$\triangleright$ EnKF:

$$
\begin{gathered}
\mathbf{x}^{a}=\mathbf{x}^{f}+\mathbf{K}\left(\mathbf{y}-H\left(\mathbf{x}^{f}\right)-\epsilon\right) \\
\mathbf{K}=\operatorname{cov}\left(\mathbf{x}_{k}, H\left(\mathbf{x}_{k}\right)\right)\left[\operatorname{cov}\left(H\left(\mathbf{x}_{k}\right)\right)+\mathbf{R}\right]^{-1}
\end{gathered}
$$

$\triangleright \mathbf{x}^{a}$ has mean and covariance matrix given by BLUE formulas
$\triangleright \mathbf{x}^{a}$ need not be Gaussian
$\triangleright$ in linear, Gaussian case, $\mathbf{x}^{a}$ has same distribution as $\mathbf{x}_{k} \mid \mathbf{y}_{1: k}$

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The EnKF is a Monte-Carlo implementation of the BLUE and, as $N_{e} \rightarrow \infty$, shares its properties.

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The EnKF is a linear method. It is optimal for linear, Gaussian systems but does not assume Gaussianity.

BLUE/EnKF Illustrated
$\triangleright \quad p(\mathbf{x})$ and ensemble


## BLUE/EnKF Illustrated

$\triangleright p(\mathbf{x} \mid y)$ for $y=x_{1}+$ noise $=1.1$ and EnKF analysis ensemble (dots)

$\triangleright$ sample retains non-Gaussian curvature but does not capture bimodality

## EnKF and Non-Gaussianity

Different EnKF schemes respond differently
$\triangleright$ All variants of EnKF produce same sample mean and 2nd moment
$\triangleright$ Other (non-Gaussian) aspects of updated ensemble depend on specific scheme
$\triangleright$ Deterministic/"square root" filters are more sensitive to nonGaussianity (Lawson and Hansen 2004, Lei et al. 2010)

Nonlinear update in observation space
$\triangleright$ EnKFs that process obs one at a time can be written as update of observed quantity followed by regression onto state variables.
$\triangleright$ Observation update is scalar and can use fully nonlinear techniques (Anderson 2010)

Iterative, Ensemble-Based Schemes
Motivation for iterations
$\triangleright$ EnKF is a linear scheme
$\triangleright$ Mean and mode of $\mathbf{x}_{k} \mid \mathbf{y}_{1: k}$ are nonlinear fns of $\mathbf{y}_{1: k}$; iteration is natural for weak nonlinearity (e.g. 4DVar)

Can EnKF be improved through iteration?
How to formulate iterations?

## Iterative, Ensemble-Based Schemes (cont.)

Several ideas
$\triangleright$ Minimize non-quadratic $J(\mathbf{x})$ with $\mathbf{x}$ restricted to ensemble subspace (Zupanski 2005)
$\triangleright$ Perform series of $N$ assimilations, each using same $\mathbf{y}_{1: k}$ but with obs-error covariance $N^{-1} \mathbf{R}$; first analysis provides prior for second, etc. (Annan et al. 2005)
$\triangleright$ Repeated application of EnKF update, mimicking the outer loop of 4DVar (Kalnay and Yang 2010)

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## 4DVar and an Iterated Ensemble Smoother

Incremental 4DVar $\equiv$ sequence of Kalman smoothers

- Linearization of $M$ and $H$ about $\mathbf{x}^{n}$ makes inner-loop $\hat{J}(\delta x)$ quadratic; thus minimization of $\hat{J}$ is equivalent to Kalman smoother
$\triangleright n$th Kalman-smoother update is

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}_{0}^{f}+\mathbf{K}_{0 \mid 1: N_{t}}\left[\mathbf{y}_{1: N_{t}}-\left(H\left(\mathbf{x}_{1: N_{t}}^{n}\right)+\mathbf{H}\left(\mathbf{x}_{1: N_{t}}^{f}-\mathbf{x}_{1: N_{t}}^{n}\right)\right)\right]
$$

$\triangleright$ see also Jazwinski (1970, section 9.7)

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$$

Approximate iterated KS using ensemble ideas
$\triangleright$ Make usual replacements:

$$
\begin{aligned}
& \mathbf{H} \delta \mathbf{x}_{k}^{f} \approx H\left(\mathbf{x}^{f}{ }_{k}\right)-H\left(\mathbf{x}_{k}^{n}\right), \\
& \mathbf{K}_{0 \mid 1: N_{t}} \approx \hat{\mathbf{K}}_{0 \mid 1: N_{t}}=\operatorname{cov}\left(\mathbf{x}_{0}, H\left(\mathbf{x}_{1: N_{t}}\right)\right)\left[\operatorname{cov}\left(H\left(\mathbf{x}_{1: N_{t}}\right)\right)+\mathbf{R}_{1: N_{t}}\right]^{-1}
\end{aligned}
$$

$\triangleright$ Ensemble ICs drawn from $N\left(\mathbf{x}_{0}^{n}, \mathbf{P}_{0}^{f}\right)$ to approximate linearization about $\mathbf{x}^{n}$ in $H$ and $M$.
$\triangleright$ Ensemble mean at iteration $n+1$ given by

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}^{f}{ }_{0}+\hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}\left(\mathbf{y}_{1: N_{t}}-\overline{H\left(\mathbf{x}_{1: N_{t} t}\right)}\right)
$$

$\triangleright$ Same as usual update, but gain changes at each iteration

## Kalnay-Yang Iteration for Ensemble KS

"Running in place" from Kalnay and Yang (2010)
$\triangleright$ Ensemble mean at iteration $n+1$ given by

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}_{0}^{n}+\hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}\left(\mathbf{y}_{1: N_{t}}-\overline{H\left(\mathbf{x}_{1: N_{t}}\right)}\right)
$$

$\triangleright$ Innovation is recalculated using most recent guess and gain changes at each iteration
$\triangleright$ Intended to speed spin up of EnKS when initial estimate of $\mathbf{P}_{0}^{f}$ is poor

Converges to observations when $H$ and $M$ are linear
$\triangleright$ Let $\mathbf{L}^{n}=\mathbf{I}-\mathbf{H}^{T} \hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}$. Easy to show

$$
\mathbf{H} \mathbf{x}_{0}^{n+1}=\left(\prod_{m=1}^{n} \mathbf{L}^{m}\right) \mathbf{H} \mathbf{x}_{0}^{f}+\left(\mathbf{I}-\prod_{m=1}^{n} \mathbf{L}^{m}\right) \mathbf{y}
$$

$\triangleright$ Properties in nonlinear case are unclear

## Simple Example: Hénon Map

## Hénon map

$\triangleright$ state is $2 \mathrm{~d}, \mathbf{x}=\left(x_{1}, x_{2}\right)$
$\triangleright$ iterate map twice in results here
$\triangleright$ Note: subscripts denote components of $\mathbf{x}$ !
An example
$\triangleright$ Gaussian ICs at $t_{0}$ ("initial time")
$\triangleright$ observe $y=x_{1}+\epsilon$ at $t_{1}$ ("final time")
$\triangleright$ update state at $t_{0}, t_{1}$

## Simple Example (cont.)

$\triangleright$ prior at $t_{1}$


## Simple Example (cont.)

$\triangleright$ prior at $t_{0}$, with value of $x_{1}\left(t_{1}\right)$ shown by colors


## Simple Example (cont.)

$\triangleright$ RMS estimation error, averaged over realizations as fn of $y$


## Particle Filters (PFs)

Sequential Monte-Carlo method to approximate $p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right)$
$\triangleright$ particles $\equiv$ ensemble members

- like EnKF, generates samples from desired pdf, rather than pdf itself


## Particle Filters (cont.)

## The simplest PF

$\triangleright \operatorname{given}\left\{\mathbf{x}_{k-1}^{i}, i=1, \ldots, N_{e}\right\}$ drawn from $p\left(\mathbf{x}_{k-1} \mid \mathbf{y}_{1: k-1}\right)$
$\triangleright \mathbf{x}_{k}^{i}=M\left(\mathbf{x}_{k-1}^{i}\right)+\epsilon_{k}$; this gives a sample from $p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k-1}\right)$.
$\triangleright$ approximate this prior as sum of point masses,

$$
p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k-1}\right) \approx N_{e}^{-1} \sum_{i=1}^{N_{e}} \delta\left(\mathbf{x}-\mathbf{x}_{k}^{i}\right)
$$

$\triangleright$ Bayes $\Rightarrow$

$$
p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right) \propto p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}\right) \sum_{i=1}^{N_{e}} \delta\left(\mathbf{x}-\mathbf{x}_{k}^{i}\right)=\sum_{i=1}^{N_{e}} p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right) \delta\left(\mathbf{x}-\mathbf{x}_{k}^{i}\right)
$$

$\triangleright$ thus, posterior pdf approximated by weighted sum of point masses

$$
p\left(\mathbf{x}_{k} \mid \mathbf{y}_{1: k}\right) \approx \sum_{i=1}^{N_{e}} w_{i} \delta\left(\mathbf{x}-\mathbf{x}_{k}^{i}\right), \quad \text { with } \quad w_{i}=\frac{p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right)}{\sum_{j=1}^{N_{e}} p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right)}
$$

## Particle Filters (cont.)

Asymptotically convergent to Bayes rule
$\triangleright$ PF yields an exact implementation of Bayes' rule as $N_{e} \rightarrow \infty$; no approximations other than finite ensemble size

Can be exceedingly simple
$\triangleright$ main calculations are for $w_{i}$, e.g. $p\left(\mathbf{y} \mid \mathbf{x}_{k}^{i}\right)$ for $i=1, \ldots, N_{e}$.
Widely applied, and effective, in low-dim'l systems
$\triangleright$ Interest for geophysical systems too: van Leeuwen $(2003,2010)$, Zhou et al. (2006), Papadakis et al. (2010), hydrology

## PF Illustrated

$\triangleright p(\mathbf{x})$, as before, and prior ensemble


## PF Illustrated

$\triangleright p(\mathbf{x} \mid \mathbf{y})$ and "weighted" ensemble (size $\propto$ weight)

$\triangleright$ weighted ensemble captures bimodality
$\triangleright$ particles don't move; assimilation is just re-weighting

## "Collapse" of Weights

A generic problem for PF
$\triangleright \max w^{i} \rightarrow 1$ as $N_{x}, N_{y}$ increase with $N_{e}$ fixed
$\triangleright$ when cycling over multiple observation times, tendency for collapse increases with $t$

## Simple Example

$\triangleright$ prior: $\mathbf{x} \sim N(0, \mathbf{I})$
$\triangleright$ identity observations: $N_{y}=N_{x}, \mathbf{H}=\mathbf{I}$
$\triangleright$ observation error: $\epsilon \sim N(0, \mathbf{I})$

Behavior of max $w^{i}$
$\triangleright \quad N_{e}=10^{3} ; N_{x}=10,30,100 ; 10^{3}$ realizations


## Required ensemble size

$\triangleright N_{e}$ s.t. PF mean has expected error less than obs


## Required ensemble size (cont.)

Collapse occurs because $w_{k}^{i}$ varies (a lot) with $i$
$\triangleright$ variance of weights (over particles, given $\mathbf{y}$ ) is controlled by

$$
\tau^{2}=\operatorname{var}\left(-\log \left(p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}\right)\right)\right)
$$

$\triangleright$ involves only obs-space quantities-no direct dependence on $N_{x}$
Conditions for collapse
$\triangleright$ if $N_{e} \rightarrow \infty$ and $\tau^{2} / \log \left(N_{e}\right) \rightarrow \infty$,

$$
E\left(1 / \max w^{i}\right) \sim 1+\frac{\sqrt{2 \log N_{e}}}{\tau}
$$

$\triangleright$ see Bengtsson et al. (2008), Snyder et al. (2008) for details
$\triangleright$ thus, weights collapse $\left(\max w^{i} \rightarrow 1\right)$ unless $N_{e}$ scales as $\exp \left(\tau^{2} / 2\right)$

## Refinements of PF

## Resampling

$\triangleright$ "refresh" ensemble by resampling from approximate posterior pdf; members with small weights are dropped, while additional members are added near members with large weights (e.g. Xiong et al. 2006, Nakano et al. 2007)
$\triangleright$ Does not overcome difficulties with PF update but reduces tendency for collapse over time

Sequential importance sampling
$\triangleright$ generate $\mathbf{x}_{k}^{i}$ using information beyond system dynamics and $\mathbf{x}_{k-1}^{i}$

## Importance Sampling

## Basic idea

$\triangleright$ Suppose $\pi(\mathbf{x})$ is hard to sample from, but $q(\mathbf{x})$ is not.
$\triangleright$ draw $\left\{\mathbf{x}^{i}\right\}$ from $q(\mathbf{x})$ and approximate

$$
\pi(\mathbf{x}) \approx \sum_{i=1}^{N_{e}} w^{i} \delta\left(\mathbf{x}-\mathbf{x}^{i}\right), \quad \text { where } w^{i}=\pi\left(\mathbf{x}^{i}\right) / q\left(\mathbf{x}^{i}\right)
$$

$\triangleright$ call $q(\mathbf{x})$ the proposal density

## Importance Sampling (cont.)

$\triangleright \quad p(\mathbf{x})$, as before, and prior ensemble

$\triangleright$ Want to sample from $p(\mathbf{x} \mid \mathbf{y})$
$\triangleright$ IS says we should weight sample from $p(\mathbf{x})$ by $p(\mathbf{x} \mid \mathbf{y}) / p(\mathbf{x})=p(\mathbf{y} \mid \mathbf{x})$

## Importance Sampling (cont.)

$\triangleright p(\mathbf{x} \mid \mathbf{y})$ and "weighted" ensemble (size $\propto$ weight)


## Sequential Importance Sampling

Perform IS sequentially in time
$\triangleright$ Given $\left\{\mathbf{x}_{0}^{i}\right\}$ from $q\left(\mathbf{x}_{0}\right)$, wish to sample from $p\left(\mathbf{x}_{1}, \mathbf{x}_{0} \mid \mathbf{y}_{1}\right)$
$\triangleright$ Note factorization:

$$
p\left(\mathbf{x}_{1}, \mathbf{x}_{0} \mid \mathbf{y}_{1}\right) \propto p\left(\mathbf{y}_{1} \mid \mathbf{x}_{1}, \mathbf{x}_{0}\right) p\left(\mathbf{x}_{1}, \mathbf{x}_{0}\right)=p\left(\mathbf{y}_{1} \mid \mathbf{x}_{1}\right) p\left(\mathbf{x}_{1} \mid \mathbf{x}_{0}\right) p\left(\mathbf{x}_{0}\right)
$$

$\triangleright$ choose proposal of the form

$$
q\left(\mathbf{x}_{1}, \mathbf{x}_{0} \mid \mathbf{y}_{1}\right)=q\left(\mathbf{x}_{1} \mid \mathbf{x}_{0}, \mathbf{y}_{1}\right) q\left(\mathbf{x}_{0}\right)
$$

$\triangleright$ update weights using

$$
w_{1}^{i} \propto \frac{p\left(\mathbf{x}_{1}^{i}, \mathbf{x}_{0}^{i} \mid \mathbf{y}_{1}\right)}{q\left(\mathbf{x}_{1}^{i}, \mathbf{x}_{0}^{i} \mid \mathbf{y}_{1}\right)}=\frac{p\left(\mathbf{y}_{1} \mid \mathbf{x}_{1}^{i}\right) p\left(\mathbf{x}_{1}^{i} \mid \mathbf{x}_{0}^{i}\right)}{q\left(\mathbf{x}_{1}^{i} \mid \mathbf{x}_{0}^{i}, \mathbf{y}_{1}\right)} w_{0}^{i}
$$

## Sequential Importance Sampling (cont.)

Choice of proposal is known to be crucial
Simplest: transition density as proposal
$\triangleright$ take $q\left(\mathbf{x}_{k} \mid \mathbf{x}_{k-1}, \mathbf{y}_{k}\right)=p\left(\mathbf{x}_{k} \mid \mathbf{x}_{k-1}\right)$; i.e. evolve particles from $t_{k-1}$ under system dynamics
$\triangleright$ weights updated by $w_{k}^{i} \propto w_{k-1}^{i} p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right)$

## Sequential Importance Sampling (cont.)

An "optimal" proposal (e.g. Doucet et al. 2000)
$\triangleright q\left(\mathbf{x}_{k} \mid \mathbf{x}_{k-1}, \mathbf{y}_{k}\right)=p\left(\mathbf{x}_{k} \mid \mathbf{x}_{k-1}, \mathbf{y}_{k}\right)$; use obs at $t_{k}$ in proposal at $t_{k}$
$\triangleright$ Papadakis et al. (2010) use this; van Leeuwen (2010) is similar
$\triangleright$ weights updated by $w_{k}^{i} \propto w_{k-1}^{i} p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k-1}^{i}\right)$
$\triangleright$ for linear, Gaussian systems, easy to show that $w_{k}^{i}$ behaves like case with prior as proposal, but $\operatorname{var}\left(\log \left(p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k-1}^{i}\right)\right)\right)$ is quantitatively smaller, by amount depending on $\mathbf{Q}$.
$N_{e}$ still grows exponentially, but w/ reduced exponent
$\triangleright$ For fixed problem, benefits can be substantial, e.g.,

$$
\begin{aligned}
& \operatorname{var}\left(\log \left(p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k-1}^{i}\right)\right)\right)=\alpha \operatorname{var}\left(\log \left(p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right)\right)\right) \Rightarrow \\
& \quad \text { ensemble size for } p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k-1}^{i}\right) \sim\left[\text { ensemble size for } p\left(\mathbf{y}_{k} \mid \mathbf{x}_{k}^{i}\right)\right]^{\alpha}
\end{aligned}
$$

## Mixture (or Gaussian-Sum) Filters

Approximate pdfs as sums of Gaussians
$\triangleright$ Start with $\left\{\mathbf{x}^{i}, \mathbf{P}^{i}\right\}$. Approximate prior pdf as

$$
p(\mathbf{x})=\sum_{i=1}^{N_{e}} w^{i} N\left(\mathbf{x} ; \mathbf{x}^{i}, \mathbf{P}^{i}\right)
$$

$\triangleright$ To compute $p(\mathbf{x} \mid \mathbf{y})$ must update $w^{i}$ (via PF-like eqns) and $\mathbf{x}^{i}, \mathbf{P}^{i}$ (via KF-like eqns); see Alspach and Sorenson (1972)
$\triangleright$ Geophysical interest: Anderson and Anderson (1999), Bengtsson et al. (2003), Smith (2007), Hoteit et al. (2011)

Limitations
$\triangleright$ Update of weights subject to collapse, as in PF; closely related to optimal proposal if we choose $\mathbf{P}^{i}=\mathbf{Q}$
$\triangleright$ Must update $\left\{\mathbf{x}^{i}, \mathbf{P}^{i}\right\}$ in addition to weights

## Summary

EnKF as approximation to BLUE
$\triangleright$ EnKF $\neq$ assume everything is Gaussian
$\triangleright$ Non-Gaussian aspects depend on specific EnKF scheme
Iterated ensemble smoother
$\triangleright$ Mimics incremental 4DVar but not equivalent (except in linear, Gaussian case!)
$\triangleright$ Innovation fixed, gain changes at each iteration
Particle filters
$\triangleright$ For naive particle filter, $N_{e}$ increases exponentially with problem size
$\triangleright$ Potential for PF using more clever proposal distributions
$\triangleright$ Evidence that these lead to $N_{e}$ that still increases exponentially, but with smaller exponent

## Comments

How important is non-Gaussianity for our applications?
A key idea missing from PFs (so far) is localization

## Selected References

## Nonlinear modifications of the EnKF:

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## 4DVar and an Iterated Kalman Smoother

## Recall 4DVar

$\triangleright$ Consider perfect model/strong constraint for simplicity here. $\mathbf{x}_{0}$ determines $\mathbf{x}_{1: N_{t}}$ through $\mathbf{x}_{k}=M\left(\mathbf{x}_{k-1}\right)$.
$\triangleright$ Full cost function from $\log \left(p\left(\mathbf{x}_{0} \mid \mathbf{y}_{1: k}\right)\right)$ :

$$
\begin{aligned}
J\left(\mathbf{x}_{0}\right)= & \left(\mathbf{x}_{0}-\mathbf{x}^{f}{ }_{0}\right)^{T}\left(\mathbf{P}_{0}^{f}\right)^{-1}\left(\mathbf{x}_{0}-\mathbf{x}^{f}{ }_{0}\right) \\
& +\left(\mathbf{y}_{1: N_{t}}-H\left(\mathbf{x}_{1: N_{t}}\right)\right)^{T} \mathbf{R}_{1: N_{t}}^{-1}\left(\mathbf{y}_{1: N_{t}}-H\left(\mathbf{x}_{1: N_{t}}\right)\right),
\end{aligned}
$$

## 4DVar and an Iterated Kalman Smoother

## Recall incremental 4DVar

$\triangleright$ Linearize about latest guess, $\mathbf{x}_{0: N_{t}}^{n}$; e.g., $H\left(\mathbf{x}_{k}\right) \approx H\left(\mathbf{x}_{k}^{n}\right)+\mathbf{H} \delta \mathbf{x}_{k}$ and $\delta \mathbf{x}_{k}=\mathbf{M}_{k-1} \delta \mathbf{x}_{k-1}$
$\triangleright$ Yields quadratic cost function for increments:

$$
\begin{aligned}
\hat{J}\left(\delta \mathbf{x}_{0}\right)= & \left(\delta \mathbf{x}_{0}-\delta \mathbf{x}_{0}^{f}\right)^{T}\left(\mathbf{P}_{0}^{f}\right)^{-1}\left(\delta \mathbf{x}_{0}-\delta \mathbf{x}_{0}^{f}\right) \\
& +\left(\delta \mathbf{y}_{1: N_{t}}-\mathbf{H} \delta \mathbf{x}_{1: N_{t}}\right)^{T} \mathbf{R}_{1: N_{t}}^{-1}\left(\delta \mathbf{y}_{1: N_{t}}-\mathbf{H} \delta \mathbf{x}_{1: N_{t}}\right),
\end{aligned}
$$

$\triangleright$ Iteration: Compute $\delta \mathbf{x}_{0}^{a}$ as minimizer of $\hat{J}$; set $\mathbf{x}_{0}^{n+1}=\mathbf{x}_{0}^{n}+\delta \mathbf{x}_{0}^{a}$; compute $\mathbf{x}_{1: N_{t}}^{n+1}$ and linearize again

## Incremental 4DVar $=$ Iterated KS

Equivalent linear, Gaussian system
$\triangleright$ Consider:

$$
\begin{aligned}
& \delta \mathbf{x}_{0} \sim N\left(\delta \mathbf{x}_{0}^{f}, \mathbf{P}_{0}^{f}\right) \\
& \delta \mathbf{x}_{k}=\mathbf{M}_{k-1} \delta \mathbf{x}_{k-1} \\
& \delta \mathbf{y}_{k}=\mathbf{H} \delta \mathbf{x}_{k}+\epsilon_{k}, \quad \epsilon_{k} \sim N\left(0, \mathbf{R}_{k}\right)
\end{aligned}
$$

$\triangleright$ Cost fn from this system is $\hat{J}\left(\delta \mathbf{x}_{0}\right)$ from incremental 4DVar
Iterated Kalman smoother
$\triangleright \delta \mathbf{x}_{0}^{a}=\arg \min \hat{J}$ can also be computed with Kalman smoother:

$$
\delta \mathbf{x}_{0}^{a}=\delta \mathbf{x}_{0}^{f}+\mathbf{K}_{0 \mid 1: N_{t}}\left(\delta \mathbf{y}_{1: N_{t}}-\mathbf{H} \delta \mathbf{x}_{1: N_{t}}^{f}\right)
$$

$\triangleright$ Thus, sequence of KS updates, with $\mathbf{M}_{k}, \mathbf{H}$ and $\mathbf{K}_{0 \mid 1: N_{t}}$ from relinearization about $\mathbf{x}_{1: N_{t}}^{n}$ at each step, reproduces incremental 4DVar
$\triangleright$ Note that initial cov of $\delta \mathbf{x}_{0}$ is $\mathbf{P}_{0}^{f}$; does not change during iteration
$\triangleright$ see also Jazwinski (1970, section 9.7)

## Iterated Ensemble KS

Approximate iterated KS using ensemble ideas
$\triangleright$ Returning to full fields, KS update becomes

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}_{0}^{f}+\mathbf{K}_{0 \mid 1: N_{t}}\left(\mathbf{y}_{1: N_{t}}-\left(H\left(\mathbf{x}_{1: N_{t}}^{n}\right)+\mathbf{H} \delta \mathbf{x}_{1: N_{t}}^{f}\right)\right)
$$

$\triangleright$ Now make usual replacements

$$
\begin{aligned}
& \mathbf{H} \delta \mathbf{x}_{k}^{f} \approx H\left(\mathbf{x}_{k}^{f}\right)-H\left(\mathbf{x}_{k}^{n}\right), \\
& \mathbf{K}_{0 \mid 1: N_{t}} \approx \hat{\mathbf{K}}_{0 \mid 1: N_{t}}=\operatorname{cov}\left(\mathbf{x}_{0}, H\left(\mathbf{x}_{1: N_{t}}\right)\right)\left[\operatorname{cov}\left(H\left(\mathbf{x}_{1: N_{t}}\right)\right)+\mathbf{R}_{1: N_{t}}\right]^{-1}
\end{aligned}
$$

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\end{aligned}
$$

Iteration for ensemble smoother
$\triangleright$ Ensemble ICs drawn from $N\left(\mathbf{x}_{0}^{n}, \mathbf{P}_{0}^{f}\right)$ to approximate linearization about $\mathbf{x}^{n}$ in $H$ and $M$.
$\triangleright$ Ensemble mean at iteration $n+1$ given by

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}^{f}{ }_{0}+\hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}\left(\mathbf{y}_{1: N_{t}}-\overline{H\left(\mathbf{x}_{1: N_{t}}\right)}\right)
$$

$\triangleright$ Same as usual update, but gain changes at each iteration

## Kalnay-Yang Iteration for Ensemble KS

"Running in place" from Kalnay and Yang (2010)
$\triangleright$ Ensemble mean at iteration $n+1$ given by

$$
\mathbf{x}_{0}^{n+1}=\mathbf{x}_{0}^{n}+\hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}\left(\mathbf{y}_{1: N_{t}}-\overline{H\left(\mathbf{x}_{1: N_{t}}\right)}\right)
$$

$\triangleright$ Innovation is recalculated using most recent guess and gain changes at each iteration
$\triangleright$ Intended to speed spin up of EnKS when initial estimate of $\mathbf{P}_{0}^{f}$ is poor

Converges to observations when $H$ and $M$ are linear
$\triangleright$ Let $\mathbf{L}^{n}=\mathbf{I}-\mathbf{H}^{T} \hat{\mathbf{K}}_{0 \mid 1: N_{t}}^{n}$. Easy to show

$$
\mathbf{H} \mathbf{x}_{0}^{n+1}=\left(\prod_{m=1}^{n} \mathbf{L}^{m}\right) \mathbf{H} \mathbf{x}_{0}^{f}+\left(\mathbf{I}-\prod_{m=1}^{n} \mathbf{L}^{m}\right) \mathbf{y}
$$

$\triangleright$ Properties in nonlinear case are unclear

